ERIC PASQUALOTTO

EFFICACY AND SAFETY OF BEXAGLIFLOZIN IN PATIENTS WITH TYPE 2 DIABETES MELLITUS: A SYSTEMATIC REVIEW AND META-ANALYSIS

Trabalho apresentado à Universidade Federal de Santa Catarina, como requisito para a conclusão do Curso de Graduação em Medicina.

Florianópolis Universidade Federal de Santa Catarina 2024

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> Florianópolis Universidade Federal de Santa Catarina

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"O sucesso nasce do querer, da determinação e persistência em se chegar a um objetivo. Mesmo não atingindo o alvo, quem busca e vence obstáculos, no mínimo fará coisas admiráveis." José de Alencar

EFFICACY AND SAFETY OF BEXAGLIFLOZIN IN PATIENTS WITH TYPE 2 DIABETES MELLITUS: A SYSTEMATIC REVIEW AND META-ANALYSIS

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ABSTRACT

Aim: To assess the efficacy of bexagliflozin in reducing glycated hemoglobin (HbA1c) and the occurrence of side effects in patients with type 2 diabetes (T2DM).

Methods: We searched PubMed, Embase, Cochrane and ClinicalTrials.gov databases for placebo-controlled, randomized clinical trials until February 15, 2023. The primary outcome was change in HbA1c. We computed weighted mean differences (WMDs) for continuous outcomes and odds ratios (ORs) for binary endpoints, with 95% confidence intervals (CIs).

Results: A total of 6 studies and 3,111 patients were included, of whom 1,951 were prescribed bexagliflozin. Compared to placebo, bexagliflozin significantly reduced HbA1c levels (WMD -0.53%; 95% CI -0.75,-0.31), fasting plasma glucose levels (WMD -1.45 mmol/L; 95% CI - 2.32,-0.57), systolic blood pressure (WMD -4.66 mmHg; 95% CI -6.41,-2.92), diastolic blood pressure (WMD -2.12 mmHg; 95% CI -3.94,-0.30), body weight (WMD -1.61 Kg; 95% CI - 2.14,-1.07), and body weight in patients with a body mass index > 25 kg/m² (WMD -2.05 Kg; 95% CI -2.78,-1.31). The proportion of patients who achieved HbA1c < 7% was higher in patients who received bexagliflozin as compared with placebo (OR 1.94; 95% CI 1.36-2.78). There were no significant differences between groups regarding side effects as hypoglycemia, genital mycotic infection, urinary tract infection, diarrhea, headache, náusea, polyuria, diabetic ketoacidosis, or all-cause mortality.

Conclusions: In this meta-analysis, the use of bexagliflozin was associated with improved clinical and laboratory measures in patients with T2DM compared to placebo, with a similar profile of side effects. These findings support the efficacy of bexagliflozin in the treatment of T2DM.

Keywords: Sodium-Glucose Transporter 2 Inhibitors; Diabetes Mellitus, Type 2; Glycated Hemoglobin.

INTRODUCTION

Glycemic control is a key aspect in the management of patients with type 2 diabetes (T2DM). This effort involves lifestyle changes, blood sugar monitoring, and use of pharmacotherapies.¹ One novel class of drugs for T2DM is the sodium-glucose cotransporter 2 inhibitors (SGLT-2i), which acts directly on the kidneys without requiring insulin secretion. The sodium-glucose cotransporter 2 (SGLT-2) is the primary glucose transporter located in the apical membrane of the proximal convoluted tubule cells and is responsible for 80-97% of renal glucose reabsorption.² Thus, SGLT-2i act decreasing the renal glucose reabsorption, leading to glycosuria and consequent glycemia reduction.^{1,3,4}

Interestingly, the benefits of SGLT-2i expand well beyond glycemic control, and include a reduction in blood pressure and body mass, irrespective of glycemic status.^{1,4,5} In addition, SGLT-2i also improve cardiovascular and renal outcomes in patients with diabetes.² Although these benefits are thought to be class-related across the spectrum of different SGLT2i, newer drugs in this class require testing to demonstrate efficacy and safety.

One example of a new highly specific and potent SGLT-2i is bexagliflozin, which has been associated with a significant reduction in hemoglobin A1c (HbA1c) and glycemic levels, according to results from small randomized controlled trials.^{6,7} However, given the small sample sizes of prior studies, a pooled analysis of trials examining bexaglifozin may bring additional insights. Therefore, we aimed to perform a meta-analysis of randomized controlled trials (RCTs) evaluating the efficacy and safety of bexagliflozin in patients with T2DM.

MATERIALS AND METHODS

Eligibility criteria

Studies that met the following eligibility criteria were included: (1) RCTs; (2) comparing bexagliflozin with placebo; (3) enrolling patients with T2DM; and (4) reporting at least one

outcome of interest. We excluded (1) overlapping populations, defined as studies with overlapping institutions and recruitment periods; and (2) non-randomized studies.

Search strategy and data extraction

We systematically searched Embase, Cochrane Central Register of Controlled Trials, ClinicalTrials.gov and PubMed databases for studies meeting the eligibility criteria and published from inception to February 15, 2023. The search strategy included the terms 'type 2 diabetes' and 'bexagliflozin'. In addition, references of systematic reviews and included studies were analyzed to verify the possibility of any other eligible studies. Two authors (E.P. and R.M.) independently extracted prespecified baseline characteristics and outcome data. Disagreements were resolved by consensus between the two authors and the senior author (E.P., R.M., and C.E.A.P) after checking the reasons of any discrepancies. Data from the longest follow-up time available in the RCTs were extracted for the analysis.

The protocol for this research was submitted to International Prospective Register of Systematic Reviews (PROSPERO) with registration number CRD42022341122. The systematic review and meta-analysis followed the recommendations of the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) statement guidelines.⁸

Endpoints and subgroup analysis

Outcomes of interest were: (1) HbA1c (%); (2) proportion of subjects achieving HbA1c < 7%; (3) fasting plasma glucose (FPG) (mmol/L); (4) systolic blood pressure (SBP) (mmHg); (5) diastolic blood pressure (DBP); (6) body weight; (7) body weight in subjects with body mass index (BMI) ≥ 25 Kg/m²; (8) hypoglycemia; (9) genital mycotic infection; (10) urinary tract

infection; (11) nausea; (12) polyuria; (13) headache; (14) diarrhea; (15) gangrene; (16) amputations; (17) diabetic ketoacidosis; (18) major adverse cardiac events (MACE); and (19) all-cause mortality.

We performed two post-hoc subgroup analyses. First, exploring the efficacy of bexagliflozin 20 mg vs. placebo, as this was the most common dose among the included studies. Second, due to heterogeneity in follow-up between studies, we also conducted an analysis restricted to follow-up time of 12 and 24 weeks for HbA1c and FPG. Finally, we also performed leave-one-out sensitivity analyses for all outcomes to ensure stability of the pooled treatment effect.

Quality assessment

Cochrane Collaboration's tool for assessing risk of bias in randomized trials (Rob 2) was used to assess quality of individual RCTs.⁹ Two independent authors (E.P. and R.M.) conducted the quality assessment. Each trial received a score of high, low, or unclear risk of bias in five domains: randomization process, deviations from the intended interventions, missing outcomes, measurement of the outcome and selection of reported results. The layout was produced by Robvis.

The overall quality of evidence was analyzed according to the Grading of Recommendation, Assessment, Development and Evaluations (GRADE) guidelines.¹⁰ The RCTs were labeled with very low, low, moderate, or high quality of evidence based on the presence of risk of bias, inconsistency of results, imprecision, publication bias, and magnitude of treatment effects.

Statistical analysis

The treatment effects for binary endpoints were compared using odds-ratio (OR), with 95% confidence intervals (CI). Continuous outcomes were evaluated using weighted mean differences (WMDs). The heterogeneity was assessed with Cochrane Q-test and I² statistics; p > 0.10 and $I^2 > 25\%$ were considered significant for heterogeneity.¹¹ We used a fixed-effect model for endpoints considered to have low heterogeneity. DerSimonian and Laird random-effects model were used in outcomes with significant heterogeneity.¹² For data handling and conversion, we used guidelines of the Cochrane Handbook for Systematic Reviews of Interventions.¹³ Review Manager 5.4 was used for statistical analysis (Nordic Cochrane Center, The Cochrane Collaboration, Copenhagen, Denmark).

RESULTS

Study selection and characteristics

The initial search yielded 63 results, as detailed in Figure 1. After removal of duplicate records and ineligible studies, 12 studies remained for full review according to prespecified criteria. Of these, 6 RCTs were included in this systematic review and meta-analysis, comprising 3,111 patients.^{14–19} A total of 1,951 (62.7%) patients received bexagliflozin, while 1,160 (37.3%) received placebo. The follow-up period ranged from 12 weeks to 168 weeks. Mean age ranged from 55.4 to 69.9 years and 1,968 (63%) patients were men. Mean BMI ranged from 28.5 to 32.8 kg/m² and mean weight ranged from 78.4 to 94.6 kg. Study and participants characteristics are summarized in Table 1.

Table 1. Baseline characteristics of included studies

Study	Follow-up	Treatment doses BG	Sample size BG/PG (%)	Age, years BG/PG (SD)	Male BG/PG (%)	BMI, kg m ⁻² BG/PG (SD)	Weight, kg BG/PG (SD)	SBP, mmHg BG/PG (SD)	Insulin use at baseline (BG/PG)	GLP-1RA use at baseline (BG/PG)
Allegretti (2019)	24 weeks	20 mg	157(50.3)/ 155(49.7)	69.3(8.36)/ 69.9(8.29)	92(58.6)/ 104(67.1)	30.29(5.9)/ 30.10(5.8)	82.90(20.5)/ 82.59(21.2)	135.9(14.2)/ 137.6(14.7)	89/91	17/13
Halvorsen (2020) [†]	12 weeks	5,10, 20 mg	76(26.0)/ 72(24.7)	59.5(10.8)/ 58.8(10.4)	50(65.8)/ 42(58.3)	28.5(5.0)/ 28.5(5.5)	78.8(16.7)/ 78.7(19.7)	127.9(14.1)/ 128.1(13.6)	NA	NA
Halvorsen (2019)	96 weeks	20 mg	145(51.2)/ 138(48.8)	56.2(10.9)/ 54.9(10.3)	67(46.2)/ 49(35.5)	29.7(5.3)/ 30.6(5.5)	78.4(17.1)/ 79.7(17.4)	127.5(13.3)/ 126.9(13.5)	NA	NA
BEST (2015)	168 weeks	20 mg	1133(66.7)/ 567(33.3)	64.4(7.9)/ 64.6(8.0)	792(69.9)/ 390(68.8)	32.80(6.1)/ 32.24(5.7)	94.59(21.9)/ 92.62(19.9)	134.2(16.2)/ 133.7(16.2)	610/292	NA
Lock (2016)	24 weeks	20 mg	138(66.7)/ 69(33.3)	55.8(10.2)/ 55.4(10.5)	66(47.8)/ 34(49.3)	32.79(5.6)/ 30.48(4.6)	90.5(20.5)/ 84.6(19.7)	131.0(14.3)/ 125.6(13.8)	0/0	0/0
Lock (2017)	24 weeks	20 mg	158(49.8)/ 159(50.2)	56.0(10.0)/ 55.6(11.2)	100(63.3)/ 94(59.1)	29.67(6.4)/ 29.99(6.3)	84.58(21.9)/ 84.44(20.9)	NA/NA	NA	NA

[†]Characteristics of patients treated with bexagliflozin 20 mg. The sample size of the intervention group is 220 patients. Abbreviations: BMI, body mass index; DBP, diastolic blood pressure; GLP-1RA, glucagon-like peptide 1 receptor agonists; SBP, systolic blood pressure; SD, standard deviation; RCT, randomized controlled trial; BG, Bexagliflozin group; PG, placebo group; NA, not available.

Figure 1. PRISMA flow diagram of study screening and selection

Pooled analysis of all studies

The primary endpoint analyzed was the reduction in HbA1c (%), which significantly improved in the bexagliflozin group when compared with placebo (WMD -0.53 %; 95% CI - 0.75, -0.31; p < 0.001; I²=84%; Figure 2A). There was a significant increase in proportion of patients who achieved HbA1c < 7% (OR 1.94; 95% CI 1.36-2.78; p<0.001; I²=0%; Figure 2B) and a significant reduction in FPG in the bexagliflozin group (WMD -1.45 mmol/L; 95% CI - 2.32, -0.57; p = 0.001; I² = 89%; Figure 2C). SBP (WMD -4.66 mmHg; 95% CI -6.41, -2.92; p < 0.001; I²=32%; Figure 2D) and DBP (WMD -2.12 mmHg; 95% CI -3.94, -0.30; p=0.02; I² = 0%; Figure 2E) were significantly reduced in the group treated with bexagliflozin. There was also a significant reduction in body weight (WMD -1.61 Kg; 95% CI -2.14, -1.07; p < 0.001; I² = 0%; Figure 3A) in the bexagliflozin group, as well as significant weight loss in patients with

BMI \geq 25 kg/m² (WMD -2.05 Kg; 95% CI -2.78, -1.31; p < 0.001; I²=72%; Figure 3B).

In safety-related outcomes, there was no significant difference between groups regarding genital mycotic infections (OR 3.11; 95% CI 0.86-11.29; p=0.08; I^2 =7%; Figure 3C), hypoglycemia (OR 0.95; 95% CI 0.80-1.14; p=0.60; I^2 =0%; Figure 3D), urinary tract infection (OR 1.04; 95% CI 0.80-1.36; p=0.75; I^2 =19%; Figure 3E), diarrhea (OR 0.75; 95% CI 0.30-1.88; p=0.54; I^2 =7%; Figure S1A), headache (OR 0.60; 95% CI 0.24-1.63; p=0.34; I^2 =0%; Figure S1B), nausea (OR 0.70; 95% CI 0.48-1.01; p=0.06; I^2 =0%; Figure S1C), polyuria (OR 1.57; 95% CI 0.83-2.99; p=0.17; I^2 =0%; Figure S1D), diabetic ketoacidosis (OR 0.44; 95% CI 0.01-17.59; p=0.66; I^2 =65%; Figure S1E), and all-cause mortality (OR 0.76; 95% CI 0.47-1.24; p=0.28; I^2 =0%; Figure S1F).

Figure 2. A. HbA1c (%). **B.** Proportion of patients who achieved HbA1c < 7%. **C.** Fasting plasma glucose (mmol/L). **D.** Systolic blood pressure (mmHg). **E.** Diastolic blood pressure (mmHg).

Figure 3. A. Body weight. B. Body weight in subjects with a BMI ≥ 25 kg/m². C. Genital mycotic infection. D. Hypoglycemia E. Urinary tract infection.

Subanalysis in selected populations

In the subgroup analysis of patients treated with the 20 mg dose of bexagliflozin vs. placebo, there was a reduction in HbA1c levels (WMD -0.55%; 95% CI -0.79, -0.31; p < 0.001; $I^2=85\%$) and a higher proportion of patients achieved HbA1c < 7% (OR 2.13; 95% CI 1.47-3.08; p < 0.001; $I^2=0\%$) in bexagliflozin-treated patients. There was a significant reduction in FPG in the intervention group (WMD -1.47 mmol/L; 95% CI -2.37, -0.57; p=0.001; $I^2=88\%$). SBP was also significantly reduced by bexagliflozin 20 mg (WMD -4.81 mmHg; 95% CI -6.23,

-3.40; p < 0.001; I²=28%), as was DBP (WMD -2.27 mmHg; 95% CI -4.36, -0.18; p=0.03; I²=0%). There was a significant reduction in body weight in patients treated with bexagliflozin 20 mg (WMD -1.75 Kg; 95% CI -2.38, -1.11; p < 0.001; I²=0%).

There was no significant difference between bexagliflozin 20 mg and placebo regarding hypoglycemia (OR 0.96; 95% CI 0.80-1.15; p=0.68; I²=0%), urinary tract infection (OR 1.04; 95% CI 0.79-1.36; p=0.78; I²=18%), diarrhea (OR 0.68; 95% CI 0.26-1.79; p=0.44; I²=23%), headache (OR 0.59; 95% CI 0.21-1.65; p=0.31; I²=0%), polyuria (OR 1.71; 95% CI 0.91-3.22; p=0.10; I²=0%), and nausea (OR 0.72; 95% CI 0.49-1.04; p=0.08; I²=0%).

In a separate subgroup analysis restricted to 12-week and 24-week follow-up for glycemic control, there was a significant reduction in HbA1c at 12 weeks (WMD -0.59 %; 95% CI -0.65, -0.53; p < 0.001; I²=0%) and 24 weeks (WMD -0.44%; 95% CI -0.55, -0.34; p <0.001; I²=30%). Furthermore, there was also a reduction in FPG both at 12 weeks (WMD -1.31 mmol/L; 95% CI -1.70, -0.92; p < 0.001; I²=47%) and 24 weeks (WMD -1.40 mmol/L; 95% CI -1.60, -1.20; p < 0.001; I²=15%).

Sensitivity analysis

We performed a leave-one-out sensitivity analysis for all outcomes. Overall, there was no change in the statistical significance of outcomes in each of the leave-one-out tests, for all outcomes. This analysis is shown for the primary endpoint of HbA1c in Table S1. There was a significant reduction in the heterogeneity between studies for the outcome of FPG with the removal of Halvorsen et al., with a reduction in $I^2=89\%$ to $I^2=36\%$.¹⁶ This is likely due to the follow-up in this study, which was substantially longer than other trials (96 weeks).¹⁶ Otherwise, there was no major shift in heterogeneity in the outcomes with the removal of each individual study in the leave-one-out analyses.

Quality and evidence assessment

Figure 4 outlines individual appraisal of each RCTs included in the meta-analysis. Overall, all studies were deemed at low risk of bias. As shown in Figure 4, the funnel plot presents a symmetrical distribution of similar-weigh studies, indicating no evidence of significant publication bias.

According to the GRADE assessment, three outcomes evaluated in this study were classified as high-quality evidence: SBP, hypoglycemia, and urinary tract infection. Seven outcomes had moderate quality of evidence: HbA1c (%), proportion of patients who achieved HbA1c < 7%, FPG, DBP, body weight, body weight in patients with a BMI $\geq 25 \text{ kg/m}^2$, and genital mycotic infection. Five outcomes were classified as having low quality of evidence: diarrhea, headache, polyuria, nausea, and all-cause mortality. Only one was considered low-quality evidence, due to the reduced number of RCTs with reported events and high heterogeneity. The main domains responsible for reducing the quality of evidence of the outcomes were: inconsistency of results, due to heterogeneity, and imprecision, due to the reduced number of RCTs with reported events. Quality assessment is detailed in Table S2.

Figure 4. A. Critical appraisal of RCTs according to the Cochrane Collaboration's tool for assessing risk of bias in randomized trials. **B.** Funnel plot analysis of the reduction in HbA1c (%) shows no evidence of publication bias.

DISCUSSION

On this systematic review and meta-analysis including 6 studies and 3,111 patients, the SGLT-2i bexagliflozin was compared with placebo in patients with T2DM. The main findings were as follows. Bexagliflozin was associated with (1) a significant reduction in HbA1c and FGP; (2) an approximate 2-fold increase in the chance of patients achieving HbA1c < 7%; (3)

a significant reduction in SBP and DBP; (4) a significant reduction in body weight; and (5) no significant increase in side effects relative to placebo.

Currently, SGLT-2i are indicated as monotherapy or adjunctive therapy in the treatment of T2DM. All SGLT-2i have shown so far similar effects regarding HbA1c reduction in this patient population. A meta-analysis by Shyangdan et al. found that patients receiving SGLT-2i achieved a greater proportion of HbA1c < 7% compared with placebo.²⁰ Furthermore, another meta-analysis comparing canagliflozin, dapagliflozin, or empagliflozin with placebo also showed a positive effect of all SGLT-2i on the reduction of HbA1c and FPG.²¹ The results of our meta-analysis extend these findings to bexagliflozin, which also was demonstrated to have a significant effect in the glycemic control, in concordance with SGLT-2i literature.^{22,23}

SGLT-2i have also consistently been shown to reduce body weight and SBP.^{22, 24,25} The significant reductions seen in body weight with SGLT-2i may relate to (1) caloric loss from urinary glucose excretion due to SGLT-2i; (2) osmotic diuresis due to glucosuria; and (3) utilization of lipid substrate by SGLT-2, causing fat loss in T2DM.²³ Our meta-analysis confirms that these findings also apply to bexagliflozin, a newer drug in this class.

Diabetes mellitus is an established risk factor for atherosclerotic cardiovascular disease and heart failure.²⁶ SGLT-2i are effective in reducing cardiovascular outcomes in this patient population, particularly heart failure events.²⁶ In the EMPA-REG trial, empagliflozin reduced the composite endpoint of death from cardiovascular causes, nonfatal myocardial infarction, and nonfatal stroke, as well as reduced deaths from any cause, as compared with placebo in patients with T2DM at high risk of cardiovascular events.²⁴ Similar findings were shown with canagliflozin in the CANVAS study.²⁵ Two MACE were reported in patients treated with bexagliflozin in our meta-analysis. Of note, we did not find a significant difference between bexagliflozin and placebo in terms of all-cause mortality. Similarly, a prior meta-analysis with fewer patients also showed no significant difference in MACE comparing bexagliflozin with placebo.²⁷ However, these findings must be interpreted with caution. Although our metaanalysis represents the largest population of bexagliflozin, the pooled number of patients in our study is still far inferior to those of the CANVAS and EMPA-REG trials.^{24,25}

Currently, a few adverse effects of this class of drugs are documented, such as hypoglycemia, genital and urinary tract infections, genital mycotic infection, allergic skin reactions, hypovolemia, dysuria, and diabetic ketoacidosis.^{28,29} This meta-analysis evaluated some of these adverse effects and the results showed no significant difference in these outcomes between groups. The relative increase in polyuria has already been documented in patients treated with bexagliflozin and is associated with the recognized volume depletion of SGLT-2i.^{6,29} The risk of hypoglycemia was not significant in previous meta-analyses with empagliflozin and dapagliflozin, similar to the results of our study, although others have shown an increase in the risk of hypoglycemia with SGLT-2i.^{1,21,22,23,25,30}

Genital mycotic infections were described as an important adverse effect of SGLT-2i, although our meta-analysis found no significant difference between bexagliflozin and placebo.^{5,31} Amputation was also described as a serious and uncommon adverse effect of SGLT-2i.²⁹ The CANVAS study showed a significantly increased risk of amputation with canagliflozin.²⁵ However, a meta-analysis evaluating exclusively amputation risk with SGLT-2i suggested no significant association.³² Only one amputation was reported in the studies included in our meta-analysis.

This study has limitations. First, there was moderate to high heterogeneity in some of the outcomes analyzed, such as HbA1c. However, we performed leave-one-out sensitivity analyses and found consistent results after removal of each study from the analysis. Second, the absence of patient-level data regarding the use of insulin or glucagon-like peptide 1 receptor agonists (GLP-1RA) precluded a subgroup analysis of patients with concomitant use of these therapies. Moreover, not all studies reported on the use of these medications. And third, albeit

this study represents the largest pooled analysis of patients treated with bexagliflozin, we remain underpowered for clinical cardiovascular endpoints.

CONCLUSION

In this meta-analysis, the use of bexagliflozin was associated with improved glycemic control, a reduction in blood pressure, and weight loss in patients with T2DM, as compared with placebo, with no significant increase in adverse events. These findings support the efficacy and safety of bexagliflozin, similar to other SGLT-2i.

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FIGURES

Figure 1.

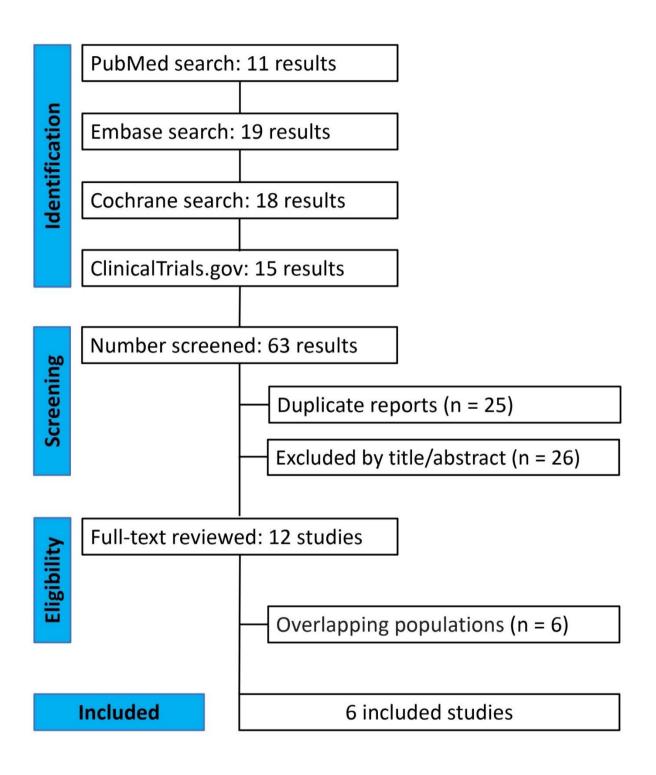
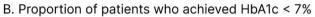


Figure 2.

A. HbA1c (%)

	Bex	in	Placebo				Mean Difference	Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Allegretti 2019	-0.59	0.814	157	-0.31	0.821	155	17.7%	-0.28 [-0.46, -0.10]	
BEST 2015	-0.56	0.828	218	-0.29	0.838	99	17.2%	-0.27 [-0.47, -0.07]	
Halvorsen 2019	-0.55	1.184	145	0.53	1.22	138	15.1%	-1.08 [-1.36, -0.80]	
Halvorsen 2020	-0.44	0.662	201	0.24	0.671	64	17.5%	-0.68 [-0.87, -0.49]	
Lock 2016	-0.51	0.946	133	-0.1	0.871	65	15.5%	-0.41 [-0.68, -0.14]	
Lock 2017	-1.09	0.906	142	-0.56	0.903	145	17.0%	-0.53 [-0.74, -0.32]	
Total (95% CI)			996			666	100.0%	-0.53 [-0.75, -0.31]	•
Heterogeneity: Tau ² =	0.06; Cł	ni² = 31.							
Test for overall effect:						,.			-1 -0.5 0 0.5 1 Favors Bexagliflozin Favors Placebo



	Bexaglif	lozin	Placebo		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Allegretti 2019	52	157	32	155	47.7%	1.90 [1.14, 3.18]	
Halvorsen 2020	58	218	11	72	26.9%	2.01 [0.99, 4.08]	
Lock 2016	41	124	13	64	25.4%	1.94 [0.95, 3.96]	
Total (95% CI)		499		291	100.0%	1.94 [1.36, 2.78]	
Total events	151		56				
Heterogeneity: Chi ² = 0	0.01, df = 2	(P = 0.9	99); l ² = 0	%			
Test for overall effect: Z = 3.62 (P = 0.0003)							Favors Placebo Favors Bexagliflozin

C. Fasting plasma glucose (mmol/L)

	Bexagliflozin F				lacebo		Mean Difference Mean Difference		
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% CI
BEST 2015	-1.06	2.486	218	-0.55	2.487	101	24.1%	-0.51 [-1.10, 0.08]	
Halvorsen 2019	-1.58	2.593	145	1.5	3.143	138	23.3%	-3.08 [-3.75, -2.41]	
Halvorsen 2020	-1.1	1.415	209	-0.11	1.412	69	25.7%	-0.99 [-1.37, -0.61]	
Lock 2016	-1.02	22.506	125	-0.15	19.84	64	1.8%	-0.87 [-7.13, 5.39]	← · · · · · · · · · · · · · · · · · · ·
Lock 2017	-2.51	2.073	142	-1.16	2.083	145	25.0%	-1.35 [-1.83, -0.87]	
Total (95% CI)			839			517	100.0%	-1.45 [-2.32, -0.57]	◆
Heterogeneity: Tau ² =	0.74; Cł	ni² = 36.3	7, df =	4 (P < 0	0.00001				
Test for overall effect: Z = 3.25 (P = 0.001)									Favors Bexagliflozin Favors Bexagliflozin

D. Systolic blood pressure (mmHg)

	Bexagliflozin Placebo roup Mean SD Total Mean SD				lacebo		Mean Difference Mean Difference				
Study or Subgroup					SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI		
BEST 2015	-9.83	15.107	417	-6.87	15.194	204	27.2%	-2.96 [-5.50, -0.42]			
Halvorsen 2019	-4.9	12.51	145	0.55	10.59	138	25.4%	-5.45 [-8.15, -2.75]			
Halvorsen 2020	-2.37	12.96	209	1.1	12.958	69	17.7%	-3.47 [-7.00, 0.06]			
Lock 2016	-0.6	41.88	138	1.54	12.659	69	4.9%	-2.14 [-9.74, 5.46]			
Lock 2017	-5.03	11.874	143	2.04	11.885	145	24.8%	-7.07 [-9.81, -4.33]			
Total (95% CI)			1052			625	100.0%	-4.66 [-6.41, -2.92]	◆		
Heterogeneity: Tau ² =	1.23; Cł	ni² = 5.87									
Test for overall effect:	Z = 5.24	(P < 0.0	0001)		-10 -5 0 5 10 Favors Bexagliflozin Favors Placebo						

E. Diastolic blood pressure (mmHg)

	Be	Bexagliflozin			Placebo			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
Halvorsen 2019	-3.07	14.471	140	-0.4	14.593	131	27.7%	-2.67 [-6.13, 0.79]	
Halvorsen 2020	-1.1	7.84	209	0.81	7.891	69	72.3%	-1.91 [-4.05, 0.23]	
Total (95% CI)			349			200	100.0%	-2.12 [-3.94, -0.30]	
Heterogeneity: Chi ² = 0.13, df = 1 (P = 0.71); l ² = 0%								-	
Test for overall effect: Z = 2.28 (P = 0.02)									Favors Bexagliflozin Favors Placebo

Figure 3.

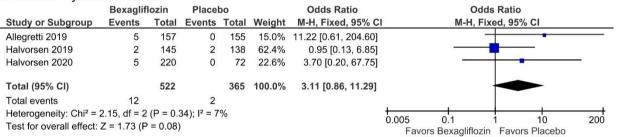
A. Body weight

	Bex	aglifloz	zin	Placebo			Mean Difference			Mean Difference			
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI		IV, Fixed	l, 95% Cl		
Halvorsen 2019	-2.95	6.279	128	-1.22	6.037	104	11.2%	-1.73 [-3.32, -0.14]		<u> </u>			
Halvorsen 2020	-1.73	2.08	209	-0.14	2.076	69	88.8%	-1.59 [-2.16, -1.02]					
Total (95% CI)			337			173	100.0%	-1.61 [-2.14, -1.07]		•			
Heterogeneity: Chi ² =	0.03, df	0.87);	$I^2 = 0\%$					+	<u> </u>		+	-+	
Test for overall effect:	Z = 5.91	I (P < 0.	00001)						-4	Favors Bexagliflozin	, Favors Pla	icebo	4

B. Body weight in subjects with a BMI $\ge 25 \text{ Kg/m}^2$

	Bexagliflozin							Mean Difference	Mean Difference		
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI		
Allegretti 2019	-2.31	2.927	122	-0.55	2.909	117	26.6%	-1.76 [-2.50, -1.02]			
BEST 2015	-3.03	3.641	922	-0.38	3.704	456	32.7%	-2.65 [-3.06, -2.24]	- -		
Lock 2016	-1.85	4.458	128	-1.06	3.819	62	18.1%	-0.79 [-2.01, 0.43]			
Lock 2017	-3.6	3.796	119	-1.09	3.741	124	22.6%	-2.51 [-3.46, -1.56]			
Total (95% CI)			1291			759	100.0%	-2.05 [-2.78, -1.31]	◆		
Heterogeneity: Tau ² =	0.38; Cł	ni² = 10.									
Test for overall effect:	Z = 5.48	8 (P < 0.	00001)						Favors Bexagliflozin Favors Placebo		

C. Genital mycotic infection



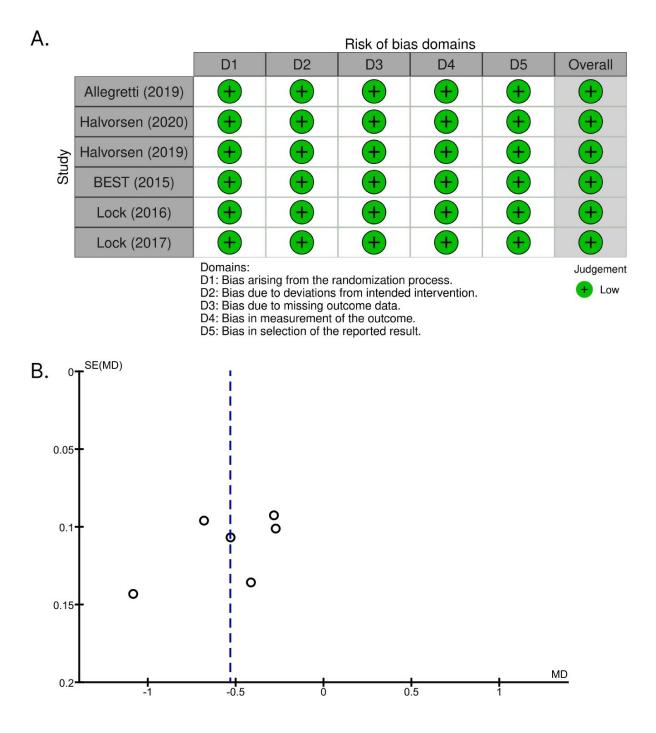
D. Hypoglycemia

	Bexaglif	lozin	Placel	bo		Odds Ratio	Odds Ratio
Study or Subgroup	Events Total		Events	s Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Allegretti 2019	39	157	38	155	11.9%	1.02 [0.61, 1.70]	
BEST 2015	475	1132	246	567	78.7%	0.94 [0.77, 1.16]	
Halvorsen 2019	24	145	25	141	8.8%	0.92 [0.50, 1.70]	
Halvorsen 2020	4	220	1	72	0.6%	1.31 [0.14, 11.96]	· · ·
Total (95% CI)		1654		935	100.0%	0.95 [0.80, 1.14]	-
Total events	542		310				
Heterogeneity: Chi ² =	0.17, df = 3	B (P = 0.9	98); I² = 0	%			
Test for overall effect:					0.5 0.7 1 1.5 2 Favors Bexagliflozin Favors Placebo		

E. Urinary tract infection

	Bexagliflozin Placebo					Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	vents Total Weigh		M-H, Fixed, 95% Cl	M-H, Fixed, 95% CI
Allegretti 2019	11	157	5	155	4.5%	2.26 [0.77, 6.67]	
BEST 2015	128	1132	60	567	67.5%	1.08 [0.78, 1.49]	
Halvorsen 2019	21	145	29	138	24.2%	0.64 [0.34, 1.18]	
Halvorsen 2020	4	220	1	72	1.4%	1.31 [0.14, 11.96]	· · · · · · · · · · · · · · · · · · ·
Lock 2016	7	138	2	69	2.4%	1.79 [0.36, 8.86]	
Total (95% CI)		1792		1001	100.0%	1.04 [0.80, 1.36]	•
Total events	171		97				
Heterogeneity: Chi ² =	4.94, df = 4	+ (P = 0.1	29); I² = 1	9%			
Test for overall effect: Z = 0.31 (P = 0.75)							0.1 0.2 0.5 1 2 5 10 Favors Bexagliflozin Favors Placebo

Figure 4.



APPENDIX

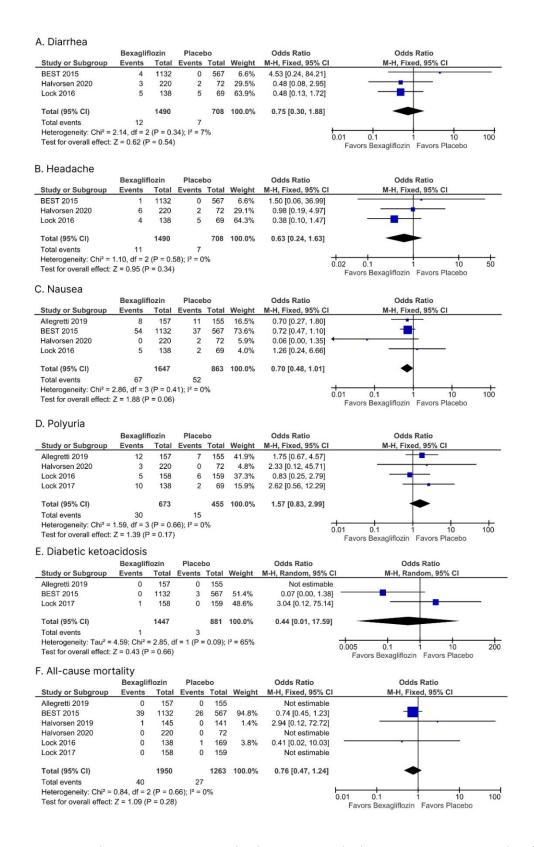


Figure S1. Forest plots outcomes. A. Diarrhea. B. Headache. C. Nausea. D. Polyuria. E.

Diabetic ketoacidosis. F. All-cause mortality.

Study removed	OR (CI 95%)	P-Value	I ²
Omitting Allegretti 2019	-0.59 (-0.83, -0.34)	P < 0.00001	84%
Omitting BEST 2015	-0.59 (-0.83, -0.34)	P < 0.00001	84%
Omitting Halvorsen 2019	-0.43 (-0.60, -0.27)	P < 0.00001	69%
Omitting Halvorsen 2020	-0.50 (-0.76, -0.25)	P < 0.00001	85%
Omitting Lock 2016	-0.56 (-0.81, -0.30)	P < 0.00001	87%
Omitting Lock 2017	-0.53 (-0.80, -0.26)	P = 0.0001	87%

 Table S1. Leave-one-out sensitivity analysis for HbA1c (%).

Table S2. Results	from the	Grading of	of Recommendations	Assessment,	Development,	and
Evaluation (GRAD	E)					

Endpoints	Patients (RCTs)	GRADE	Rationale for GRADE
HbA1c (%)	1662 (6)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, there was high heterogeneity $(I^2 = 84\%)$.
Proportion of patients who achieved HbA1c < 7%	790 (3)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, we downgraded by one level due to only 3 RCTs reporting this outcome.
FPG (mmol/L)	1356 (5)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, there was high heterogeneity ($I^2 = 89\%$).
SBP (mmHg)	1677 (5)	⊕ ⊕ ⊕ ⊕ Strong	All RCTs had a low risk of bias. However, there was significant heterogeneity ($I^2 = 32\%$), possibly due to the large reduction in SBP found in some studies compared to other RCTs or due to follow-up. Based on this, we consider that this should not reduce the strength of the evidence.
DBP (mmHg)	549 (2)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, only two studies reported this outcome.
Body weight	510 (2)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, only two studies reported this outcome.

Body weight in patients with a $BMI \ge 25 \text{ kg/m2}$	2050 (4)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, there was high heterogeneity ($I^2 = 72\%$).
Genital mycotic infection	887 (3)	$\oplus \oplus \oplus \ominus$ Moderate	All RCTs had a low risk of bias. However, only three studies reported this outcome.
Hypoglycemia	2589 (4)	$ \bigoplus \bigoplus \bigoplus \bigoplus \bigoplus $ Strong	All RCTs had a low risk of bias.
Urinary tract infection	2793 (5)	⊕ ⊕ ⊕ ⊕ Strong	All RCTs had a low risk of bias. Although the confidence intervals include evidence of no effect, the incidence was similar between the bexagliflozin 20 mg and placebo groups across the 5 RCTs. Based on this, we consider that this should not reduce the strength of the evidence.
Diarrhea	2198 (3)	$ \bigoplus \bigoplus \ominus \ominus \Theta $ Low	All RCTs had a low risk of bias. However, we downgraded by two levels due to only 3 RCTs reporting this outcome and given that the confidence intervals include evidence of no effect.
Headache	2198 (3)	$ \bigoplus \bigoplus \ominus \ominus \Theta $ Low	All RCTs had a low risk of bias. However, we downgraded by two levels due to only 3 RCTs reporting this outcome and given that the confidence intervals include evidence of no effect.
Nausea	2510 (4)	$ \bigoplus \bigoplus \ominus \ominus \ominus $ Low	All RCTs had a low risk of bias. However, we downgraded by two levels due to only 4 RCTs reporting this outcome.
Polyuria	1128 (4)	$\bigoplus \bigoplus \ominus \ominus$ Low	All RCTs had a low risk of bias. However, we downgraded by two levels due to only 4 RCTs reporting this outcome and given that the confidence intervals include evidence of no effect.
Diabetic ketoacidosis	2328 (3)	$\bigoplus \ominus \ominus \ominus$ Very Low	All RCTs had a low risk of bias. However, we downgraded by three levels due to only 3 RCTs reporting this outcome and given that the confidence intervals include evidence of no effect. Furthermore, only 2 RCTs that reported events were evaluated with meta-analysis, with high heterogeneity ($I^2 = 65\%$).
All-cause mortality	3213 (6)	$\bigoplus \bigoplus \ominus \ominus$ Low	All RCTs had a low risk of bias. However, we downgraded by two levels due to only 3 RCTs reporting events and given that the confidence intervals include evidence of no effect.

BMI, body mass index; DBP, diastolic blood pressure; FPG, fasting plasma glucose; GRADE, Grading of Recommendations Assessment, Development, and Evaluation; HbA1c, hemoglobin A1c; RCTs, randomized controlled trials; SBP, systolic blood pressure.

ANEXO 1: Normas de publicação da revista Diabetes, Obesity and Metabolism

AIMS AND SCOPE

Authors are reminded that *Diabetes, Obesity and Metabolism* is primarily a journal of pharmacology and therapeutics, focused mainly on human research relevant to patient care. Priority is given to manuscripts that relate to therapeutic interventions.

The scope of the journal includes studies of pharmacokinetics and pharmacodynamics; costeffectiveness; real world evidence of drug utilisation, safety and effectiveness, as well as conventional randomised controlled trials. High-quality meta-analyses and systematic reviews that provide original information on treatment effects and safety are considered as original research papers.

DOM also welcomes manuscripts which report clinical data relating to novel devices, Apps, glucose sensors and insulin pumps used to improve glycaemic control.

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Authors should note that submission of a manuscript implies that the content has not been published previously (except in abstract form, e.g as part of conference proceedings), and is not currently under consideration by any other journal.

Once the submission materials have been prepared in accordance with the guidance for authors, manuscripts should be submitted online at: <u>http://mc.manuscriptcentral.com/dom</u>

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Diabetes Obesity and Metabolism is participating in a scheme to improve Peer Review Transparency. We encourage all authors to opt in to transparent peer review (TPR) at the time of manuscript submission.

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Allowing readers access to the peer review process and pre-acceptance clarifications and revisions adds value and invariably reflects positively on the authors, sponsors and the journal. It is pleasing that >90% of authors are currently opting into TPR, but if authors have any concerns they are welcome to raise these with the handling editor.

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MANUSCRIPT CATEGORIES AND REQUIREMENTS

(i) Original Papers

Original research papers should provide substantial new analyses and be structured as follows:

Title page: Author names, affiliations, and a short running title. The number of co-authors should not exceed 15. The word counts for the abstract and the main body of the text (excluding references and legends) should be clearly stated, along with the number of references, tables and figures. Please conform to the guidance on manuscript size below

Structured abstract: Subheadings - aims, materials and methods, results, and conclusions (maximum 250 words);

Main manuscript: This should be typed double-spaced and structured as follows: introduction; materials and methods (including appropriate subsections, e.g. statistical methods); results; discussion; acknowledgements; references (normally <60); legends to figures; tables; and figures. It is helpful to include line numbering throughout the document.

Original manuscripts should not exceed **3,500 words** (not including references, tables and figures). Manuscripts should include a maximum of 5 Figures and/or tables.

Additional tables or figures and/or extra methodological detail can be included in a separate Supplementary Appendix. The production and handling Editors may relocate large tables or figures into a Supplementary Appendix prior to the production of page proofs if the manuscript exceeds these specifications at the time of final acceptance.

(ii) Research Letters

These are short, focused communications conveying original results from new research which is more limited in scope and depth than what would be expected for an original paper. Nevertheless, Research Letters undergo full peer review; they are fully searchable and citeable items with their own unique DOI number.

Research Letters are ideal to communicate a focused piece of research, e.g a post-hoc or secondary analysis of a large clinical trial, a limited meta-analysis, or a specific new result from real world evidence. Originality and clinical relevance are important.

Research Letters <u>do not</u> have an abstract but they should have a title page (as above).

The main text should be sub-divided into sections, e.g : Background / context, Methods, Results and Conclusions.

The total manuscript length should not exceed **1,200 words**, excluding references.

Research Letters can include a maximum of 12 references (+ any number of references to work published in *Diab. Obes. Metab.*) and 2 figures OR 2 tables (or one of each).

Additional information (e.g text, tables or figures) can be provided as part of a Supplementary Appendix.

Research Letters have now replaced Brief Reports.

(iii) Review Articles & Commentaries

Diabetes, Obesity & Metabolism publishes a limited number of narrative reviews. (Systematic reviews and meta-analyses should be formatted and submitted as original papers).

We consider unsolicited review articles on important topics of therapeutics and pharmacology. Good review articles should provide novel insights, informed discussion and a balanced interpretation of the published literature.

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The editors welcome correspondence relating to work that has recently been published in the journal, and/or other brief comments or observations that may be of wider interest.

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We do not publish case reports.

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Following the acceptance and online publication of an article in *Diab Obes. Metab.*, authors may wish to submit one or more digital enhancements (DE's), e.g a video abstract, infographic or plain language summary (PLS), to augment the communication.

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1. King VM, Armstrong DM, Apps R, Trott JR. Numerical aspects of pontine, lateral reticular, and inferior olivary projections to two paravermal cortical zones of the cat cerebellum. J Comp Neurol 1998;390:537-551.

Book

1. Voet D, Voet JG. Biochemistry. New York: John Wiley & Sons; 1990. 1223 p.

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